

MAS651 Theory of Stochastic Processes

Homework #5

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Throughout this homework, let \mathbb{Z}_+ denote the set of all non-negative integers, \mathbb{R}_+ be the set of all non-negative real numbers, and $[a : b] := \{a, a + 1, \dots, b - 1, b\}$ for $a, b \in \mathbb{Z}$ with $a \leq b$. We also write $[n] := [1 : n]$ for $n \in \mathbb{N}$. Moreover, \uplus denotes the *disjoint union*, and given a set A and $k \in \mathbb{Z}_+$, $\binom{A}{k} := \{B \subseteq A : |B| = k\}$. Also, we use the symbol \mathbb{S} instead of S to denote the underlying state space of stochastic processes.

We assume throughout this homework that the underlying state space \mathbb{S} is countable and it is equipped with the discrete σ -field $2^{\mathbb{S}}$ on \mathbb{S} . Since $(\mathbb{S}, 2^{\mathbb{S}})$ is a *nice measurable space*, it admits the canonical construction in *Section 5.2* in [1] of the probability measure \mathbb{P}_μ on the sequence space $(\Omega_0, \mathcal{F}_\infty)$ so that the sequence of coordinate maps $\{X_n(\omega) := \omega_n\}_{n=0}^\infty$ is a homogeneous Markov chain with initial distribution μ and transition probability $p(\cdot, \cdot) : \mathbb{S} \times 2^{\mathbb{S}} \rightarrow [0, 1]$. We remark that it is conventional to write $p(x, y) := p(x, \{y\})$ for $x, y \in \mathbb{S}$.

Problem 1 (*Exercise 5.6.1* in [1]).

Observe that for any $n \in \mathbb{Z}_+$, we have $\mathbb{1}_{\{X_{n+1}=0\}} = \mathbb{1}_{\{X_1=0\}} \circ \theta_n$ on Ω_0 . Thus,

$$\begin{aligned} \mathbb{P}_\mu \{X_{n+1} = 0\} &= \mathbb{E}_\mu [\mathbb{1}_{\{X_{n+1}=0\}}] \\ &= \mathbb{E}_\mu [\mathbb{1}_{\{X_1=0\}} \circ \theta_n] \\ &= \mathbb{E}_\mu [\mathbb{E}_\mu [\mathbb{1}_{\{X_1=0\}} \circ \theta_n | \mathcal{F}_n]] \\ &\stackrel{(a)}{=} \mathbb{E}_\mu [\mathbb{E}_{X_n} [\mathbb{1}_{\{X_1=0\}}]] \\ &= \mathbb{E}_\mu \left[\sum_{x \in \mathbb{S}} \mathbb{E}_{X_n} [\mathbb{1}_{\{X_1=0\}}] \mathbb{1}_{\{X_n=x\}} \right] \\ &= \mathbb{P}_0 \{X_1 = 0\} \cdot \mathbb{P}_\mu \{X_n = 0\} + \mathbb{P}_1 \{X_1 = 0\} \cdot \mathbb{P}_\mu \{X_n = 1\} \\ &= (1 - \alpha) \mathbb{P}_\mu \{X_n = 0\} + \beta (1 - \mathbb{P}_\mu \{X_n = 0\}) \\ &= \beta + (1 - \alpha - \beta) \mathbb{P}_\mu \{X_n = 0\}, \end{aligned} \tag{1}$$

where the step (a) holds by the Markov property (*Theorem 5.2.3* in [1]). Here, $\{\mathcal{F}_n := \sigma(X_0, X_1, \dots, X_n)\}_{n=0}^\infty$ denotes the canonical filtration of the Markov chain $\{X_n\}_{n=0}^\infty$. From (1), we obtain

$$\mathbb{P}_\mu \{X_n = 0\} - \frac{\beta}{\alpha + \beta} = (1 - \alpha - \beta) \left(\mathbb{P}_\mu \{X_{n-1} = 0\} - \frac{\beta}{\alpha + \beta} \right) \tag{2}$$

for all $n \in \mathbb{N}$. Now, we deduce our desired result via induction on n . The case $n = 0$ is trivial. Now, assume that we have

$$\mathbb{P}_\mu \{X_{k-1} = 0\} = \frac{\beta}{\alpha + \beta} + (1 - \alpha - \beta)^{k-1} \left\{ \mu(0) - \frac{\beta}{\alpha + \beta} \right\} \quad (3)$$

for $k \geq 1$. Putting the induction hypothesis (3) into the recursive relation (2) directly yields

$$\mathbb{P}_\mu \{X_k = 0\} = \frac{\beta}{\alpha + \beta} + (1 - \alpha - \beta)^k \left\{ \mu(0) - \frac{\beta}{\alpha + \beta} \right\},$$

which completes the proof.

Problem 2 (*Exercise 5.6.2. in [1]*).

Since $p(\cdot, \cdot) : \mathbb{S} \times \mathbb{S} \rightarrow [0, 1]$ is irreducible, there is a positive integer $K(x, y) \in \mathbb{N}$ such that $p^{K(x,y)}(x, y) > 0$. Due to the aperiodicity of $p(\cdot, \cdot)$, we have $d_x = 1$ and so there is a positive integer $N(x) \in \mathbb{N}$ such that $p^n(x, x) > 0$ for all $n \geq N(x)$ by *Lemma 5.6.5 in [1]*. Thus, one has

$$p^{n+N(x)+K(x,y)}(x, y) \stackrel{(a)}{\geq} p^{n+N(x)}(x, x) \cdot p^{K(x,y)}(x, y) > 0 \quad (4)$$

for all $n \in \mathbb{Z}_+$, where the step (a) follows from the Chapman-Kolmogorov equation. Let

$$\Gamma := \max \{N(x) + K(x, y) : (x, y) \in \mathbb{S} \times \mathbb{S}\},$$

which is finite since $\mathbb{S} \times \mathbb{S}$ is finite. Then for any $n \geq \Gamma$,

$$p^n(x, y) > 0, \quad \forall (x, y) \in \mathbb{S} \times \mathbb{S},$$

since $n \geq \Gamma \geq N(x) + K(x, y)$ for all $(x, y) \in \mathbb{S} \times \mathbb{S}$ together with the fact (4). Hence, we have $p^n(x, y) > 0$ for all $(x, y) \in \mathbb{S} \times \mathbb{S}$, for any $n \geq \Gamma$, and it suffices to choose the desired integer m to be greater than Γ .

Remark 1. Let $(\mathbb{S}, \mathcal{S})$ be a nice state space, and $(\Omega_0, \mathcal{F}_\infty)$ denote the sequence space obtained from $(\mathbb{S}, \mathcal{S})$. For any probability measure $\mu : \mathcal{S} \rightarrow [0, 1]$ on $(\mathbb{S}, \mathcal{S})$, let \mathbb{P}_μ denote the canonical probability measure on the sequence space $(\Omega_0, \mathcal{F}_\infty)$ constructed in *Section 5.2 of [1]* via Kolmogorov's extension theorem, and $\mathbb{P}_x := \mathbb{P}_{\delta_x}$ for $x \in \mathbb{S}$, where $\delta_x : \mathcal{S} \rightarrow [0, 1]$ refers to the Dirac measure centered on the state $x \in \mathbb{S}$. Then, we know that

$$\mathbb{P}_\mu \{E\} = \int_{\mathbb{S}} \mu(dx) \mathbb{P}_x \{E\} \quad (5)$$

for all $E \in \mathcal{F}_\infty$.

Problem 3 (*Exercise 5.6.3. in [1]*).

To begin with, we note that $\mathbb{S}^2 := \mathbb{S} \times \mathbb{S}$ is finite.

Claim 1. *The transition probability $\bar{p}(\cdot, \cdot) : \mathbb{S}^2 \times \mathbb{S}^2 \rightarrow [0, 1]$ on \mathbb{S}^2 defined by*

$$\bar{p}((x_1, y_1), (x_2, y_2)) := p(x_1, x_2) \cdot p(y_1, y_2), \quad \forall (x_1, y_1), (x_2, y_2) \in \mathbb{S}^2,$$

is irreducible and aperiodic.

Proof of Claim 1.

We first claim that for all $n \in \mathbb{N}$,

$$\bar{p}^n((x_1, y_1), (x_2, y_2)) \geq p^n(x_1, x_2) \cdot p^n(y_1, y_2), \quad \forall (x_1, y_1), (x_2, y_2) \in \mathbb{S}^2. \quad (6)$$

The proof of the claim (6) hinges upon the induction on n . The case $n = 1$ is immediate from the definition of $\bar{p}(\cdot, \cdot)$. Now, assume that (6) holds for the case $n = k - 1$, where $k \geq 2$. Then,

$$\begin{aligned} \bar{p}^k((x_1, y_1), (x_2, y_2)) &\stackrel{(a)}{=} \sum_{(z, w) \in \mathbb{S}^2} \bar{p}((x_1, y_1), (z, w)) \bar{p}^{k-1}((z, w), (x_2, y_2)) \\ &\stackrel{(b)}{\geq} \sum_{(z, w) \in \mathbb{S}^2} p(x_1, z) p(y_1, w) \cdot p^{k-1}(z, x_2) p^{k-1}(w, y_2) \\ &= \left\{ \sum_{z \in \mathbb{S}} p(x_1, z) p^{k-1}(z, x_2) \right\} \left\{ \sum_{w \in \mathbb{S}} p(y_1, w) p^{k-1}(w, y_2) \right\} \\ &\stackrel{(c)}{=} p^k(x_1, x_2) p^k(y_1, y_2), \end{aligned}$$

and this proves the claim (6). Here, the above steps (a)–(c) can be justified as follows:

- (a) the Chapman-Kolmogorov equation;
- (b) the induction hypothesis;
- (c) the same reason as the step (a).

Due to the irreducibility of $p(\cdot, \cdot)$, there are positive integers $K, L \in \mathbb{N}$ such that

$$p^K(x_1, x_2) > 0 \quad \text{and} \quad p^L(y_1, y_2) > 0.$$

Also from the aperiodicity of $p(\cdot, \cdot)$, there is a positive integer $N_0 \in \mathbb{N}$ such that

$$p^n(x_1, x_1) > 0 \quad \text{and} \quad p^n(y_1, y_1) > 0$$

for all $n \geq N_0$. Thus, we have for every $n \geq N_0$,

$$\begin{aligned} p^{L+n+K}(x_1, x_2) &\stackrel{(d)}{\geq} p^{L+n}(x_1, x_1) \cdot p^K(x_1, x_2) > 0; \\ p^{K+n+L}(y_1, y_2) &\stackrel{(e)}{\geq} p^{K+n}(y_1, y_1) \cdot p^L(y_1, y_2) > 0, \end{aligned} \quad (7)$$

where the step (d) and (e) are consequences of the Chapman-Kolmogorov equation. So for all $n \geq K + L + N_0$, one has

$$\bar{p}^n((x_1, y_1), (x_2, y_2)) \stackrel{(f)}{\geq} p^n(x_1, x_2) \cdot p^n(y_1, y_2) \stackrel{(g)}{>} 0, \quad (8)$$

where the step (f) is due to Claim 1, and the step (g) is due to (7), and this establishes the irreducibility of $\bar{p}(\cdot, \cdot)$. Note that the integer $K + L + N_0$ depends on the choice of two states $(x_1, y_1), (x_2, y_2) \in \mathbb{S}^2$. Moreover, putting $(x_1, y_1) = (x_2, y_2) = (x, y) \in \mathbb{S}^2$ into the inequality (8) yields $\bar{p}^n((x, y), (x, y)) > 0$ for all but finitely many $n \in \mathbb{N}$. This implies $d_{(x, y)} = 1$ for all $(x, y) \in \mathbb{S}^2$, thereby $\bar{p}(\cdot, \cdot)$ is aperiodic. \square

Combining Claim 1 together with Problem 2 guarantees that there exists a positive integer $\Gamma \in \mathbb{N}$ such that

$$\bar{p}^\Gamma((x_1, y_1), (x_2, y_2)) > 0, \quad \forall (x_1, y_1), (x_2, y_2) \in \mathbb{S}^2. \quad (9)$$

Let $\{Z_n := (X_n, Y_n)\}_{n=0}^\infty$ be the canonical homogeneous Markov chain constructed via the construction on the sequence space in Section 5.2 of [1] with state space \mathbb{S}^2 and transition probability $\bar{p}(\cdot, \cdot) : \mathbb{S}^2 \times \mathbb{S}^2 \rightarrow [0, 1]$. Further we let $\Delta := \{(x, x) \in \mathbb{S}^2 : x \in \mathbb{S}\}$ denote the *diagonal of \mathbb{S}^2* , and

$$\epsilon := \min \left\{ \sum_{z \in \mathbb{S}} \bar{p}^\Gamma((x, y), (z, z)) : (x, y) \in \mathbb{S}^2 \right\} = \min \{ \mathbb{P}_{(x,y)} \{Z_\Gamma \in \Delta\} : (x, y) \in \mathbb{S}^2 \} > 0.$$

At this point, recall that $T := \inf \{n \geq 1 : Z_n \in \Delta\}$. Then, we have

$$\begin{aligned} \mathbb{P}_{(x,y)} \{T > \Gamma\} &= \mathbb{P}_{(x,y)} \{Z_1 \in \mathbb{S}^2 \setminus \Delta, Z_2 \in \mathbb{S}^2 \setminus \Delta, \dots, Z_\Gamma \in \mathbb{S}^2 \setminus \Delta\} \\ &\leq \mathbb{P}_{(x,y)} \{Z_\Gamma \in \mathbb{S}^2 \setminus \Delta\} \\ &= 1 - \mathbb{P}_{(x,y)} \{Z_\Gamma \in \Delta\} \\ &\leq 1 - \epsilon \end{aligned} \quad (10)$$

for all $(x, y) \in \mathbb{S}^2$. Thanks to Remark 1, we arrive at

$$\begin{aligned} \mathbb{P}_\nu \{T > \Gamma\} &= \sum_{(x,y) \in \mathbb{S}^2} \nu(x, y) \cdot \mathbb{P}_{(x,y)} \{T > \Gamma\} \\ &\leq (1 - \epsilon) \sum_{(x,y) \in \mathbb{S}^2} \nu(x, y) \\ &= 1 - \epsilon, \end{aligned} \quad (11)$$

where $\nu(\cdot) : \mathbb{S}^2 \rightarrow [0, 1]$ is any initial distribution $\{Z_n\}_{n=0}^\infty$. One can observe that for each $k \geq 2$, we have that if $T(\omega) > (k-1)\Gamma$,

$$\begin{aligned} \mathbb{1}_{\{T > \Gamma\}} \circ \theta_{(k-1)\Gamma}(\omega) &= \begin{cases} 1 & \text{if } Z_1(\theta_{(k-1)\Gamma}(\omega)) \in \mathbb{S}^2 \setminus \Delta, \dots, Z_\Gamma(\theta_{(k-1)\Gamma}(\omega)) \in \mathbb{S}^2 \setminus \Delta; \\ 0 & \text{otherwise} \end{cases} \\ &= \begin{cases} 1 & \text{if } Z_{(k-1)\Gamma+1}(\omega) \in \mathbb{S}^2 \setminus \Delta, \dots, Z_{k\Gamma}(\omega) \in \mathbb{S}^2 \setminus \Delta; \\ 0 & \text{otherwise} \end{cases} \\ &\stackrel{\text{(h)}}{=} \begin{cases} 1 & \text{if } Z_1(\omega) \in \mathbb{S}^2 \setminus \Delta, \dots, Z_{k\Gamma}(\omega) \in \mathbb{S}^2 \setminus \Delta; \\ 0 & \text{otherwise} \end{cases} \\ &= \mathbb{1}_{\{T > k\Gamma\}}(\omega), \end{aligned}$$

where the step (h) holds since $T(\omega) > (k-1)\Gamma$. In other words,

$$\mathbb{1}_{\{T > k\Gamma\}} = (\mathbb{1}_{\{T > \Gamma\}} \circ \theta_{(k-1)\Gamma}) \mathbb{1}_{\{T > (k-1)\Gamma\}} \quad (12)$$

on $(\mathbb{S}^2)^{\mathbb{Z}_+}$. Here, $\theta_n : (\mathbb{S}^2)^{\mathbb{Z}_+} \rightarrow (\mathbb{S}^2)^{\mathbb{Z}_+}$ denotes the shift operator on $(\mathbb{S}^2)^{\mathbb{Z}_+}$. Therefore,

$$\begin{aligned}
\mathbb{P}_\nu \{T > k\Gamma\} &= \mathbb{E}_\nu [\mathbb{1}_{\{T > k\Gamma\}}] \\
&\stackrel{(i)}{=} \mathbb{E}_\nu [(\mathbb{1}_{\{T > \Gamma\}} \circ \theta_{(k-1)\Gamma}) \mathbb{1}_{\{T > (k-1)\Gamma\}}] \\
&= \mathbb{E}_\nu \left[\mathbb{E}_\nu \left[(\mathbb{1}_{\{T > \Gamma\}} \circ \theta_{(k-1)\Gamma}) \mathbb{1}_{\{T > (k-1)\Gamma\}} \mid \mathcal{F}_{(k-1)\Gamma}^{\mathbf{Z}} \right] \right] \\
&\stackrel{(j)}{=} \mathbb{E}_\nu \left[\mathbb{E}_\nu \left[\mathbb{1}_{\{T > \Gamma\}} \circ \theta_{(k-1)\Gamma} \mid \mathcal{F}_{(k-1)\Gamma}^{\mathbf{Z}} \right] \mathbb{1}_{\{T > (k-1)\Gamma\}} \right] \\
&\stackrel{(k)}{=} \mathbb{E}_\nu \left[\mathbb{E}_{Z_{(k-1)\Gamma}} [\mathbb{1}_{\{T > \Gamma\}}] \mathbb{1}_{\{T > (k-1)\Gamma\}} \right] \\
&\stackrel{(l)}{\leq} (1 - \epsilon) \mathbb{P}_\nu \{T > (k-1)\Gamma\},
\end{aligned} \tag{13}$$

where the above steps (i)–(l) can be validated as follows:

(i) the equality (12);

(j) $\{T > (k-1)\Gamma\} = (\mathbb{S}^2)^{\mathbb{Z}_+} \setminus \{T \leq (k-1)\Gamma\} \in \mathcal{F}_{(k-1)\Gamma}^{\mathbf{Z}}$, since T is a stopping time with respect to the canonical filtration $\{\mathcal{F}_n^{\mathbf{Z}}\}_{n=0}^\infty$ denotes the canonical filtration of the Markov chain $\{Z_n\}_{n=0}^\infty$, where

$$\mathcal{F}_n^{\mathbf{Z}} := \sigma(Z_0, Z_1, \dots, Z_n) = \sigma((X_0, Y_0), (X_1, Y_1), \dots, (X_n, Y_n))$$

for each $n \in \mathbb{Z}_+$;

(k) the Markov property (*Theorem 5.2.3* in [1]);

(l) the inequality (11).

So, we may inductively deduce from (13) that

$$\mathbb{P}_\nu \{T > k\Gamma\} \leq (1 - \epsilon)^k \tag{14}$$

for every $k \in \mathbb{Z}_+$. Note that the bound (14) holds for any initial distribution $\nu(\cdot) : \mathbb{S}^2 \rightarrow [0, 1]$ of the Markov chain $\{Z_n\}_{n=0}^\infty$.

Finally, choose any $n \in \mathbb{Z}_+$ and let $k := \lfloor \frac{n}{\Gamma} \rfloor \in \mathbb{Z}_+$. Since $k\Gamma \leq n < (k+1)\Gamma$, we know $\frac{n}{\Gamma} - 1 < k \leq \frac{n}{\Gamma}$. Hence,

$$\begin{aligned}
\mathbb{P}_\nu \{T > n\} &\leq \mathbb{P}_\nu \{T > k\Gamma\} \\
&\stackrel{(m)}{\leq} (1 - \epsilon)^k \\
&\stackrel{(n)}{\leq} (1 - \epsilon)^{\frac{n}{\Gamma} - 1} \\
&= \frac{1}{1 - \epsilon} \cdot \left\{ (1 - \epsilon)^{\frac{1}{\Gamma}} \right\}^n,
\end{aligned} \tag{15}$$

where the step (m) makes use of the bound (14), and the step (n) holds since $0 < 1 - \epsilon < 1$ and $k > \frac{n}{\Gamma} - 1$. By letting $C := \frac{1}{1 - \epsilon} \in (0, +\infty)$ and $r := (1 - \epsilon)^{\frac{1}{\Gamma}} \in (0, 1)$, the bound (15) establishes the desired result.

Problem 4 (*Exercise 5.6.5* in [1]: Strong law for additive functionals).

(i) We first prove the following useful result inspired by *Exercise 5.3.1* in [1]:

Lemma 1. Let $\{X_n\}_{n=0}^\infty$ be a homogeneous Markov chain with countable state space \mathbb{S} and transition probability $p(\cdot, \cdot) : \mathbb{S} \times \mathbb{S} \rightarrow [0, 1]$, and $x \in \mathbb{S}$ be a recurrent state of the chain. For $k \in \mathbb{N}$, let $\tau_k := T_x^k - T_x^{k-1}$ be the k -th inter-arrival time to state x , and $V_k := \left(\tau_k, X_{T_x^{k-1}}, X_{T_x^{k-1}+1}, \dots, X_{T_x^k-1} \right)$. Then given any probability distribution $\mu(\cdot) : \mathbb{S} \rightarrow [0, 1]$, the sequence of random vectors $\{V_k : k \geq 2\}$ are independent and identically distributed under the canonical probability measure \mathbb{P}_μ defined on the sequence space $(\Omega_0, \mathcal{F}_\infty)$, conditionally on the event $\{T_x < +\infty\}$, where $T_x := T_x^1$ is the first hitting time to state x .

Proof of Lemma 1.

To begin with, we may observe that if $T_x^{k-1}(\omega) < +\infty$, then $\mathbb{1}_{\{V_1=v\}} \circ \theta_{T_x^{k-1}} = \mathbb{1}_{\{V_k=v\}}$ for all $k \geq 2$ and $v \in \mathbb{V} := \bigcup_{n=1}^\infty (\{n\} \times \mathbb{S}^n)$. That is,

$$\mathbb{1}_{\{V_k=v\}} = \left(\mathbb{1}_{\{V_1=v\}} \circ \theta_{T_x^{k-1}} \right) \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \quad (16)$$

on Ω_0 , since $\{V_k = v\} \subseteq \{T_x^k < +\infty\} \subseteq \{T_x^{k-1} < +\infty\}$. Thus, for every $k \geq 2$, one has

$$\begin{aligned} \mathbb{P}_\mu \left\{ V_k = v \mid \mathcal{F}_{T_x^{k-1}} \right\} &= \mathbb{E}_\mu \left[\mathbb{1}_{\{V_k=v\}} \mid \mathcal{F}_{T_x^{k-1}} \right] \\ &\stackrel{(a)}{=} \mathbb{E}_\mu \left[\left(\mathbb{1}_{\{V_1=v\}} \circ \theta_{T_x^{k-1}} \right) \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \mid \mathcal{F}_{T_x^{k-1}} \right] \\ &\stackrel{(b)}{=} \mathbb{E}_\mu \left[\mathbb{1}_{\{V_1=v\}} \circ \theta_{T_x^{k-1}} \mid \mathcal{F}_{T_x^{k-1}} \right] \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \\ &\stackrel{(c)}{=} \mathbb{E}_{X_{T_x^{k-1}}} \left[\mathbb{1}_{\{V_1=v\}} \right] \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \\ &\stackrel{(d)}{=} \mathbb{E}_x \left[\mathbb{1}_{\{V_1=v\}} \right] \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \\ &= \mathbb{P}_x \{V_1 = v\} \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \end{aligned} \quad (17)$$

\mathbb{P}_μ -almost surely, where the above steps (a)–(d) can be justified as follows:

(a) the equality (16);

(b) $\{T_x^{k-1} < +\infty\} \in \mathcal{F}_{T_x^{k-1}}$, since

$$\{T_x^{k-1} < +\infty\} \cap \{T_x^{k-1} = n\} = \{T_x^{k-1} = n\} \in \mathcal{F}_{T_x^{k-1}}$$

for every $n \in \mathbb{Z}_+$;

(c) the strong Markov property (*Theorem 5.2.5* in [1]);

(d) if $T_x^{k-1} < +\infty$, then $X_{T_x^{k-1}} = x$ for $k \geq 2$.

One can immediately deduce from (17) that

$$\mathbb{P}_\mu \{V_k = v\} = \mathbb{E}_\mu \left[\mathbb{P}_\mu \left\{ V_k = v \mid \mathcal{F}_{T_x^{k-1}} \right\} \right] = \mathbb{P}_x \{V_1 = v\} \mathbb{P}_\mu \left\{ T_x^{k-1} < +\infty \right\}. \quad (18)$$

At this point, we claim that for all $k \in \mathbb{N}$, $\mathbb{P}_\mu \{T_x^k < +\infty\} = \mathbb{P}_\mu \{T_x < +\infty\}$. If $k = 1$, there's nothing to

prove and we may assume that $k \geq 2$. One can easily see that if $T_x^{k-1}(\omega) < +\infty$, then

$$\begin{aligned} \left(\mathbb{1}_{\{T_x < +\infty\}} \circ \theta_{T_x^{k-1}} \right) (\omega) &= \begin{cases} 1 & \text{if } X_n \left(\theta_{T_x^{k-1}}(\omega) \right) = x \text{ for some } n > 0; \\ 0 & \text{otherwise} \end{cases} \\ &= \begin{cases} 1 & \text{if } X_n(\omega) = x \text{ for some } n > T_x^{k-1}(\omega); \\ 0 & \text{otherwise} \end{cases} \\ &= \mathbb{1}_{\{T_x^k < +\infty\}}(\omega). \end{aligned}$$

More succinctly, we have

$$\mathbb{1}_{\{T_x^k < +\infty\}} = \left(\mathbb{1}_{\{T_x < +\infty\}} \circ \theta_{T_x^{k-1}} \right) \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \quad (19)$$

on Ω_0 . Hence,

$$\begin{aligned} \mathbb{P}_\mu \left\{ T_x^k < +\infty \right\} &= \mathbb{E}_\mu \left[\mathbb{1}_{\{T_x^k < +\infty\}} \right] \\ &\stackrel{(e)}{=} \mathbb{E}_\mu \left[\left(\mathbb{1}_{\{T_x < +\infty\}} \circ \theta_{T_x^{k-1}} \right) \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \right] \\ &= \mathbb{E}_\mu \left[\mathbb{E}_\mu \left[\left(\mathbb{1}_{\{T_x < +\infty\}} \circ \theta_{T_x^{k-1}} \right) \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \middle| \mathcal{F}_{T_x^{k-1}} \right] \right] \\ &\stackrel{(f)}{=} \left[\mathbb{E}_\mu \left[\mathbb{1}_{\{T_x < +\infty\}} \circ \theta_{T_x^{k-1}} \middle| \mathcal{F}_{T_x^{k-1}} \right] \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \right] \\ &\stackrel{(g)}{=} \mathbb{E}_\mu \left[\mathbb{E}_{X_{T_x^{k-1}}} \left[\mathbb{1}_{\{T_x < +\infty\}} \right] \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \right] \\ &\stackrel{(h)}{=} \mathbb{E}_\mu \left[\mathbb{E}_x \left[\mathbb{1}_{\{T_x < +\infty\}} \right] \mathbb{1}_{\{T_x^{k-1} < +\infty\}} \right] \\ &= \rho_{xx} \cdot \mathbb{P}_\mu \left\{ T_x^{k-1} < +\infty \right\} \\ &\stackrel{(i)}{=} \mathbb{P}_\mu \left\{ T_x^{k-1} < +\infty \right\}, \end{aligned} \quad (20)$$

where the above steps (e)–(i) hold since:

- (e) the equality (19);
- (f) the same reason as the step (b);
- (g) the same reason as the step (c);
- (h) the same reason as the step (d);
- (i) $\rho_{xx} = 1$, because the state $x \in \mathbb{S}$ is recurrent,

thereby one can deduce our desired claim from (20) inductively. So, the equation (18) becomes

$$\mathbb{P}_x \{V_1 = v\} \mathbb{P}_\mu \{T_x < +\infty\} = \mathbb{P}_\mu \{V_k = v\} \stackrel{(j)}{=} \mathbb{P}_\mu \{V_k = v, T_x < +\infty\},$$

where the step (j) holds since $\{V_k = v\} \subseteq \{T_x < +\infty\}$, thereby we arrive at

$$\mathbb{P}_x \{V_1 = v\} = \mathbb{P}_\mu \{V_k = v | T_x < +\infty\} \quad (21)$$

for all $k \geq 2$. Thus, $\{V_k : k \geq 2\}$ are identically distributed under the canonical probability measure \mathbb{P}_μ on the sequence space $(\Omega_0, \mathcal{F}_\infty)$, conditionally on the event $\{T_x < +\infty\}$.

As the final step, it remains to establish the conditional independence of the sequence $\{V_k : k \geq 2\}$ given the event $\{T_x < +\infty\}$. Let $\{v_k : k \geq 2\}$ be any sequence in \mathbb{V} . Since $\mathbb{P}_\mu\{T_x < +\infty\} = \mathbb{P}_\mu\{T_x^k < +\infty\}$ for all $k \in \mathbb{N}$, we know that $\mathbb{1}_{\mathbb{P}_\mu\{T_x < +\infty\}} \stackrel{\mathbb{P}_\mu\text{-a.s.}}{=} \mathbb{1}_{\mathbb{P}_\mu\{T_x^k < +\infty\}}$ for all $k \in \mathbb{N}$. Let us begin our argument from the equation (17): for every $k \geq 2$,

$$\mathbb{P}_\mu \left\{ V_k = v_k \mid \mathcal{F}_{T_x^{k-1}} \right\} \stackrel{\mathbb{P}_\mu\text{-a.s.}}{=} \mathbb{P}_x \{V_1 = v_k\} \mathbb{1}_{\{T_x < +\infty\}}. \quad (22)$$

It is easy to see that $\{V_2 = v_2, V_3 = v_3, \dots, V_{k-1} = v_{k-1}\} \in \mathcal{F}_{T_x^{k-1}}$, so

$$\begin{aligned} & \mathbb{P}_\mu \{V_2 = v_2, V_3 = v_3, \dots, V_k = v_k\} \\ &= \int_{\{V_2=v_2, V_3=v_3, \dots, V_{k-1}=v_{k-1}\}} \mathbb{1}_{\{V_k=v_k\}} d\mathbb{P}_\mu \\ &\stackrel{(k)}{=} \int_{\{V_2=v_2, V_3=v_3, \dots, V_{k-1}=v_{k-1}\}} \mathbb{P}_x \{V_1 = v_k\} \mathbb{1}_{\{T_x < +\infty\}} d\mathbb{P}_\mu \\ &= \mathbb{P}_x \{V_1 = v_k\} \mathbb{P}_\mu \{V_2 = v_2, V_3 = v_3, \dots, V_{k-1} = v_{k-1}, T_x < +\infty\} \\ &\stackrel{(l)}{=} \mathbb{P}_\mu \{V_k = v_k \mid T_x < +\infty\} \mathbb{P}_\mu \{V_2 = v_2, V_3 = v_3, \dots, V_{k-1} = v_{k-1} \mid T_x < +\infty\} \mathbb{P}_\mu \{T_x < +\infty\}, \end{aligned} \quad (23)$$

where the step (k) comes from (22), and the step (l) is due to the equation (21). Hence, we reach

$$\begin{aligned} & \mathbb{P}_\mu \{V_2 = v_2, V_3 = v_3, \dots, V_{k-1} = v_{k-1} \mid T_x < +\infty\} \mathbb{P}_\mu \{V_k = v_k \mid T_x < +\infty\} \\ &= \frac{\mathbb{P}_\mu \{V_2 = v_2, V_3 = v_3, \dots, V_k = v_k\}}{\mathbb{P}_\mu \{T_x < +\infty\}} \\ &\stackrel{(m)}{=} \frac{\mathbb{P}_\mu \{V_2 = v_2, V_3 = v_3, \dots, V_k = v_k, T_x < +\infty\}}{\mathbb{P}_\mu \{T_x < +\infty\}} \\ &= \mathbb{P}_\mu \{V_2 = v_2, V_3 = v_3, \dots, V_k = v_k \mid T_x < +\infty\}, \end{aligned} \quad (24)$$

where the step (m) holds since $\{V_2 = v_2, V_3 = v_3, \dots, V_k = v_k\} \subseteq \{T_x < +\infty\}$. Therefore, we may inductively conclude that

$$\mathbb{P}_\mu \{V_2 = v_2, V_3 = v_3, \dots, V_n = v_n \mid T_x < +\infty\} = \prod_{k=2}^n \mathbb{P}_\mu \{V_k = v_k \mid T_x < +\infty\}$$

for all $n \geq 2$ and $v_2, v_3, \dots, v_n \in \mathbb{V}$. Hence, $\{V_k : k \geq 2\}$ is a sequence of independent and identically distributed random vectors under \mathbb{P}_μ defined on $(\Omega_0, \mathcal{F}_\infty)$, conditionally on the event $\{T_x < +\infty\}$. \square

Now, it's time to complete the proof of the statement (i). Since $p(\cdot, \cdot)$ is irreducible and has a stationary distribution $\pi(\cdot) : \mathbb{S} \rightarrow [0, 1]$, it is positive recurrent by *Theorem 5.5.12* in [1]. Thanks to Lemma 1, for any state $x \in \mathbb{S}$, the sequence of random vectors,

$$\left\{ V_k := \left(\tau_k, X_{T_x^{k-1}}, X_{T_x^{k-1}+1}, \dots, X_{T_x^k-1} \right) : k \geq 2 \right\},$$

are independent and identically distributed under \mathbb{P}_μ , conditionally on the event $\{T_x < +\infty\}$. Here, we may observe that for any event $E \in \mathcal{F}_\infty$,

$$\begin{aligned} \mathbb{P}_\mu \{E\} &= \mathbb{P}_\mu \{E \mid T_x < +\infty\} \mathbb{P}_\mu \{T_x < +\infty\} + \mathbb{P}_\mu \{E \mid T_x = +\infty\} \mathbb{P}_\mu \{T_x = +\infty\} \\ &\stackrel{(n)}{=} \mathbb{P}_\mu \{E \mid T_x < +\infty\}, \end{aligned} \quad (25)$$

where the step (n) can be justified as follows: since $p(\cdot, \cdot)$ is irreducible and the state $x \in \mathbb{S}$ is recurrent, $\rho_{yx} = 1$ for all $y \in \mathbb{S}$ due to *Theorem 5.3.2* in [1]. So, Remark 1 implies

$$\mathbb{P}_\mu \{T_x < +\infty\} = \sum_{y \in \mathbb{S}} \mu(y) \cdot \mathbb{P}_y \{T_x < +\infty\} = \sum_{y \in \mathbb{S}} \mu(y) \cdot \rho_{yx} = 1,$$

as desired. From the observation (25), one can see from Lemma 1 that $\{V_k\}_{k=2}^\infty$ is a sequence of independent and identically distributed random vectors under the canonical probability measure \mathbb{P}_μ defined on the sequence space $(\Omega_0, \mathcal{F}_\infty)$, not necessarily conditionally on the event $\{T_x < +\infty\}$. Furthermore, we obtain from (21) together with the observation (25) that

$$\mathbb{P}_x \{V_1 = v\} = \mathbb{P}_\mu \{V_k = v\} \tag{26}$$

for all $k \geq 2$ and $v \in \mathbb{V}$. Now, define the function $F : \mathbb{V} \rightarrow \mathbb{R}$ by

$$F(n, x_0, x_1, \dots, x_{n-1}) := \sum_{j=0}^{n-1} f(x_j), \quad \forall (n, x_0, x_1, \dots, x_{n-1}) \in \mathbb{V}.$$

Since $V_k^f = F(V_{k+1})$ for $k \in \mathbb{N}$, $\{V_k^f\}_{k=1}^\infty$ is a sequence of independent and identically distributed random variables under \mathbb{P}_μ , for any initial distribution $\mu(\cdot) : \mathbb{S} \rightarrow [0, 1]$ of the Markov chain.

It still remains to prove the \mathbb{P}_π -integrability of V_k^f for $k \geq 2$, i.e., $\mathbb{E}_\pi \left[\left| V_k^f \right| \right] < +\infty$ for all $k \geq 2$. To begin with, we introduce the following conventions:

$$\pi(f) := \sum_{y \in \mathbb{S}} f(y) \pi(y) \quad \text{and} \quad \pi(|f|) := \sum_{y \in \mathbb{S}} |f(y)| \pi(y).$$

Then, the following bound holds:

$$\begin{aligned}
\mathbb{E}_\pi \left[\left| V_k^f \right| \right] &= \mathbb{E}_\pi \left[\left| \sum_{j=T_x^k}^{T_x^{k+1}-1} f(X_j) \right| \right] \\
&\leq \mathbb{E}_\pi \left[\sum_{j=T_x^k}^{T_x^{k+1}-1} |f(X_j)| \right] \\
&\stackrel{(o)}{=} \sum_{1 \leq a < b < +\infty} \mathbb{E}_\pi \left[\sum_{j=T_x^k}^{T_x^{k+1}-1} |f(X_j)| \middle| T_x^k = a, T_x^{k+1} = b \right] \mathbb{P}_\pi \left\{ T_x^k = a, T_x^{k+1} = b \right\} \\
&= \sum_{1 \leq a < b < +\infty} \mathbb{E}_\pi \left[\sum_{j=a}^{b-1} |f(X_j)| \right] \mathbb{P}_\pi \left\{ T_x^k = a, T_x^{k+1} = b \right\} \\
&= \sum_{1 \leq a < b < +\infty} \left[\sum_{j=a}^{b-1} \underbrace{\mathbb{E}_\pi [|f(X_j)|]}_{= \pi(|f|)} \right] \mathbb{P}_\pi \left\{ T_x^k = a, T_x^{k+1} = b \right\} \\
&\stackrel{(p)}{=} \pi(|f|) \sum_{j=a}^{b-1} (b-a) \mathbb{P}_\pi \left\{ T_x^k = a, T_x^{k+1} = b \right\} \\
&= \pi(|f|) \sum_{n=1}^{\infty} n \cdot \mathbb{P}_\pi \left\{ \tau_{k+1} = n \right\} \\
&\stackrel{(q)}{=} \pi(|f|) \sum_{n=1}^{\infty} n \cdot \mathbb{P}_x \left\{ T_x = n \right\} \\
&= \pi(|f|) \mathbb{E}_x [T_x] \stackrel{(r)}{<} +\infty,
\end{aligned}$$

where the above steps (o)–(r) can be validated as follows:

(o) we have $\mathbb{P}_\pi \left\{ T_x^k < +\infty \right\} = 1$ for all $n \in \mathbb{N}$ due to Remark 1;

(p) $\pi(\cdot) : \mathbb{S} \rightarrow [0, 1]$ is a stationary distribution for $p(\cdot, \cdot)$;

(q) from (26), we get

$$\mathbb{P}_x \left\{ T_x = n \right\} = \mathbb{P}_\mu \left\{ \tau_k = n \right\}$$

for all $n \in \mathbb{N}$, $k \geq 2$, and any initial distribution $\mu(\cdot)$ of the Markov chain;

(r) the state $x \in \mathbb{S}$ is positive recurrent,

and this completes the proof of the statement (i).

(ii) Firstly, one can see that

$$\begin{aligned}
K_n &:= \inf \left\{ k \in \mathbb{Z}_+ : T_x^k \geq n \right\} \\
&= \inf \left\{ k \in \mathbb{Z}_+ : T_x^k > n - 1 \right\} \\
&= \inf \left\{ k \in \mathbb{Z}_+ : \sum_{j=1}^{n-1} \mathbb{1}_{\{X_j=x\}} \leq k - 1 \right\} \\
&= \sum_{j=1}^{n-1} \mathbb{1}_{\{X_j=x\}} + 1 \\
&= N_{n-1}(x) + 1.
\end{aligned} \tag{27}$$

We defined the events

$$\begin{aligned}
\mathcal{E}_1 &:= \left\{ \omega \in \Omega_0 : \lim_{n \rightarrow \infty} \frac{K_n(\omega)}{n} = \frac{1}{\mathbb{E}_x[T_x]} \right\} \in \mathcal{F}_\infty; \\
\mathcal{E}_2 &:= \left\{ \omega \in \Omega_0 : \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n V_k^f(\omega) = \mathbb{E}_\pi[V_1^f] \right\} \in \mathcal{F}_\infty.
\end{aligned}$$

Due to *Theorem 5.6.1* in [1], we know that for any state $y \in \mathbb{S}$,

$$\frac{N_n(x)}{n} \xrightarrow{n \rightarrow \infty} \frac{1}{\mathbb{E}_x[T_x]} \mathbb{1}_{\{T_x < +\infty\}} \stackrel{(r)}{=} \frac{1}{\mathbb{E}_x[T_x]}$$

\mathbb{P}_y -almost surely, where the step (r) follows from the fact $\mathbb{P}_y\{T_x < +\infty\} = \rho_{yx} = 1$ for every $y \in \mathbb{S}$, which holds by *Theorem 5.3.2* in [1]. Thus,

$$\frac{K_n}{n} \stackrel{(s)}{=} \frac{N_{n-1}}{n-1} \left(1 - \frac{1}{n}\right) + \frac{1}{n} \xrightarrow{n \rightarrow \infty} \frac{1}{\mathbb{E}_x[T_x]}$$

\mathbb{P}_y -almost surely, where the step (s) makes use of the observation (27). So, we get $\mathbb{P}_y\{\mathcal{E}_1\} = 1$ for all $y \in \mathbb{S}$.

On the other hand, we know from the statement (i) that $\mathbb{E}_\pi[V_k^f] < +\infty$ for all $k \geq 2$. The strong law of large number yields $\mathbb{P}_\pi\{\mathcal{E}_2\} = 1$. Since the transition probability $p(\cdot, \cdot)$ is irreducible, there is a positive integer $K(x, y) \in \mathbb{N}$ such that $p^{K(x,y)}(x, y) > 0$ for every $(x, y) \in \mathbb{S} \times \mathbb{S}$. If there exists a state $z \in \mathbb{S}$ such that $\pi(z) = 0$, then

$$0 = \pi(z) = \sum_{x \in \mathbb{S}} \pi(x) \cdot p^{K(y,z)}(x, z) \geq \pi(y) \cdot p^{K(y,z)}(y, z),$$

which implies $\pi(y) = 0$ for all $y \in \mathbb{S}$. This violates the fact that $\pi(\cdot)$ has total mass 1, *i.e.*, $\sum_{x \in \mathbb{S}} \pi(x) = 1$, and therefore we find that $\pi(x) > 0$ for all $x \in \mathbb{S}$. Remark 1 implies

$$1 = \mathbb{P}_\pi\{\mathcal{E}_2\} = \sum_{y \in \mathbb{S}} \pi(y) \mathbb{P}_y\{\mathcal{E}_2\},$$

and this yields $\mathbb{P}_y\{\mathcal{E}_2\} = 1$ for all $y \in \mathbb{S}$, since $\pi(y) > 0$ for all $y \in \mathbb{S}$. Hence, we arrive at $\mathbb{P}_y\{\mathcal{E}_1 \cap \mathcal{E}_2\} = 1$ for all $y \in \mathbb{S}$, and employing Remark 1 again yields

$$\mathbb{P}_\mu\{\mathcal{E}_1 \cap \mathcal{E}_2\} = \sum_{y \in \mathbb{S}} \mu(y) \mathbb{P}_y\{\mathcal{E}_1 \cap \mathcal{E}_2\} = \sum_{y \in \mathbb{S}} \mu(y) = 1$$

for any initial distribution $\mu(\cdot) : \mathbb{S} \rightarrow [0, 1]$ of the Markov chain.

Finally, for every $\omega \in \mathcal{E}_1 \cap \mathcal{E}_2$, one can see that $\lim_{n \rightarrow \infty} K_n(\omega) = +\infty$ since $\mathbb{E}_x [T_x] < +\infty$. Thus,

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^{K_n(\omega)} V_k^f(\omega) &= \lim_{n \rightarrow \infty} \frac{K_n(\omega)}{n} \cdot \frac{1}{K_n(\omega)} \sum_{k=1}^{K_n(\omega)} V_k^f(\omega) \\ &= \frac{\mathbb{E}_\pi [V_1^f]}{\mathbb{E}_x [T_x]} \\ &\stackrel{(t)}{=} \pi(f), \end{aligned}$$

where the step (t) can be verified as follows:

$$\begin{aligned} \mathbb{E}_\pi [V_1^f] &= \mathbb{E}_\pi \left[\sum_{j=T_x^k}^{T_x^{k+1}-1} f(X_j) \right] \\ &\stackrel{(u)}{=} \sum_{1 \leq a < b < +\infty} \mathbb{E}_\pi \left[\sum_{j=T_x^k}^{T_x^{k+1}-1} f(X_j) \middle| T_x^k = a, T_x^{k+1} = b \right] \mathbb{P}_\pi \{T_x^k = a, T_x^{k+1} = b\} \\ &= \sum_{1 \leq a < b < +\infty} \mathbb{E}_\pi \left[\sum_{j=a}^{b-1} f(X_j) \right] \mathbb{P}_\pi \{T_x^k = a, T_x^{k+1} = b\} \\ &= \sum_{1 \leq a < b < +\infty} \left[\sum_{j=a}^{b-1} \underbrace{\mathbb{E}_\pi [f(X_j)]}_{=\pi(f)} \right] \mathbb{P}_\pi \{T_x^k = a, T_x^{k+1} = b\} \\ &\stackrel{(v)}{=} \pi(f) \sum_{j=a}^{b-1} (b-a) \mathbb{P}_\pi \{T_x^k = a, T_x^{k+1} = b\} \\ &= \pi(f) \sum_{n=1}^{\infty} n \cdot \mathbb{P}_\pi \{\tau_{k+1} = n\} \\ &\stackrel{(w)}{=} \pi(f) \sum_{n=1}^{\infty} n \cdot \mathbb{P}_x \{T_x = n\} \\ &= \pi(f) \mathbb{E}_x [T_x], \end{aligned}$$

where the steps (u), (v), and (w) hold by the same reason as the steps (o), (p), and (q), respectively. So,

$$\frac{1}{n} \sum_{m=1}^{K_n} V_m^f \xrightarrow{n \rightarrow \infty} \frac{\mathbb{E}_\pi [V_1^f]}{\mathbb{E}_x [T_x]} = \pi(f)$$

\mathbb{P}_μ -almost surely, for any initial distribution $\mu(\cdot)$ of the Markov chain.

(iii) To begin with, we provide the following critical lemma:

Lemma 2. *Let $\{X_n\}_{n=1}^{\infty}$ be a sequence of independent and identically distributed random variables defined on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$ with $\mathbb{E}[|X_1|] < +\infty$. Then,*

$$\frac{1}{n} \max \{|X_k| : k \in [n]\} \xrightarrow{\mathbb{P}\text{-a.s.}} 0 \tag{28}$$

as $n \rightarrow \infty$.

Proof of Lemma 2.

Firstly, we fix any $\epsilon > 0$. Then,

$$\begin{aligned}
\sum_{n=1}^{\infty} \mathbb{P}\{|X_n| \geq n\epsilon\} &\stackrel{(x)}{=} \sum_{n=1}^{\infty} \mathbb{P}\{|X_1| \geq n\epsilon\} \\
&= \sum_{n=1}^{\infty} \int_{n-1}^n \mathbb{P}\left\{\frac{|X_1|}{\epsilon} \geq n\right\} dt \\
&\stackrel{(y)}{\leq} \sum_{n=1}^{\infty} \int_{n-1}^n \mathbb{P}\left\{\frac{|X_1|}{\epsilon} \geq t\right\} dt \\
&= \int_0^{\infty} \mathbb{P}\left\{\frac{|X_1|}{\epsilon} \geq t\right\} dt \\
&\stackrel{(z)}{=} \frac{\mathbb{E}[|X_1|]}{\epsilon} < +\infty,
\end{aligned} \tag{29}$$

where the above steps (x)–(z) can be verified as follows:

- (x) the sequence $\{X_n\}_{n=1}^{\infty}$ are identically distributed;
- (y) for $n-1 \leq t \leq n$, we have $\mathbb{P}\left\{\frac{|X_1|}{\epsilon} \geq n\right\} \leq \mathbb{P}\left\{\frac{|X_1|}{\epsilon} \geq t\right\}$;
- (z) Lemma 2.2.13 in [1].

From the observation (29), we employ the first Borel-Cantelli lemma:

$$\mathbb{P}\left\{\limsup_{n \rightarrow \infty} \{|X_n| \geq n\epsilon\}\right\} = 0.$$

So, we conclude that for any $\epsilon > 0$,

$$\mathbb{P}\left\{\liminf_{n \rightarrow \infty} \{|X_n| < n\epsilon\}\right\} = 1. \tag{30}$$

Now, let $A_k := \liminf_{n \rightarrow \infty} \left\{\frac{|X_n|}{n} < \frac{1}{k}\right\}$ for $k \in \mathbb{N}$, and $A := \bigcap_{k=1}^{\infty} A_k$. From (30), we know $\mathbb{P}\{A\} = 1$ and if $\omega \in A$, then $\frac{|X_n(\omega)|}{n} < \frac{1}{k}$ for all but finitely many $n \in \mathbb{N}$. Thus,

$$\limsup_{n \rightarrow \infty} \frac{|X_n(\omega)|}{n} \leq \frac{1}{k}$$

for all $k \in \mathbb{N}$, and letting $k \rightarrow \infty$ yields the desired result. □

By replacing f by $|f|$ in the statement (i), one can see that $\{V_k^{|f|} : k \in \mathbb{N}\}$ is a sequence of independent and identically distributed random variables under the canonical probability measure \mathbb{P}_μ on the sequence space $(\Omega_0, \mathcal{F}_\infty)$, for any initial distribution $\mu(\cdot)$ of the Markov chain, and $\mathbb{E}_\pi[V_k^{|f|}] < +\infty$ for all $k \in \mathbb{N}$. Applying Lemma 2 gives

$$\frac{1}{n} \max\{V_k^{|f|} : k \in [n]\} \xrightarrow{n \rightarrow \infty} 0 \tag{31}$$

\mathbb{P}_π -almost surely. Let

$$\begin{aligned}
\mathcal{E}_3 &:= \left\{\omega \in \Omega_0 : \lim_{n \rightarrow \infty} \frac{1}{n} \max\{V_k^{|f|}(\omega) : k \in [n]\} = 0\right\} \in \mathcal{F}_\infty; \\
\mathcal{E}_4 &:= \left\{\omega \in \Omega_0 : T_x^k(\omega) < +\infty \text{ for all } k \in \mathbb{N}\right\} \in \mathcal{F}_\infty.
\end{aligned}$$

Then, (31) implies $\mathbb{P}_\pi \{\mathcal{E}_3\} = 1 = \sum_{y \in \mathbb{S}} \pi(y) \cdot \mathbb{P}_y \{\mathcal{E}_3\}$. Since $\pi(y) > 0$ for all $y \in \mathbb{S}$, we have $\mathbb{P}_y \{\mathcal{E}_3\} = 1$ for all $y \in \mathbb{S}$. Hence, Remark 1 yields

$$\mathbb{P}_\mu \{\mathcal{E}_3\} = \sum_{y \in \mathbb{S}} \mu(y) \cdot \mathbb{P}_y \{\mathcal{E}_3\} = \sum_{y \in \mathbb{S}} \mu(y) = 1,$$

and this establishes $\frac{1}{n} \max \left\{ V_k^{|f|} : k \in [n] \right\} \xrightarrow{n \rightarrow \infty} 0$, \mathbb{P}_μ -almost surely, for any initial distribution $\mu(\cdot)$ of the Markov chain.

Furthermore, the irreducibility of $p(\cdot, \cdot)$ together with the recurrence of the state $x \in \mathbb{S}$ yields $\mathbb{P}_y \{\mathcal{E}_4\} = 1$ for all $y \in \mathbb{S}$. Therefore, we arrive at $\mathbb{P}_y \{\mathcal{E}_3 \cap \mathcal{E}_4\} = 1$ for all $y \in \mathbb{S}$, thereby from Remark 1,

$$\mathbb{P}_\mu \{\mathcal{E}_3 \cap \mathcal{E}_4\} = \sum_{y \in \mathbb{S}} \mu(y) \cdot \mathbb{P}_y \{\mathcal{E}_3 \cap \mathcal{E}_4\} = \sum_{y \in \mathbb{S}} \mu(y) = 1$$

for any initial distribution $\mu(\cdot)$ of the Markov chain. Since we know that $\mathbb{P}_\mu \{\mathcal{E}_1 \cap \mathcal{E}_2\}$, we finally obtain

$$\mathbb{P}_\mu \{\mathcal{E}_1 \cap \mathcal{E}_2 \cap \mathcal{E}_3 \cap \mathcal{E}_4\} = 1 \quad (32)$$

for any initial distribution $\mu(\cdot) : \mathbb{S} \rightarrow [0, 1]$ of the Markov chain. At this point, we propose the following decomposition:

$$\begin{aligned} \sum_{j=1}^n f(X_j) &= \sum_{j=1}^{T_x-1} f(X_j) + \sum_{m=1}^{K_n-2} \left[\sum_{j=T_x^m}^{T_x^{m+1}-1} f(X_j) \right] + \sum_{j=T_x^{K_n-1}}^n f(X_j) \\ &= \sum_{j=1}^{T_x-1} f(X_j) + \sum_{m=1}^{K_n-2} V_m^f + \sum_{j=T_x^{K_n-1}}^n f(X_j), \end{aligned}$$

which leads to

$$\frac{1}{K_n} \sum_{j=1}^n f(X_j) = \underbrace{\frac{1}{K_n} \sum_{j=1}^{T_x-1} f(X_j)}_{(T1)} + \underbrace{\left(1 - \frac{2}{K_n}\right) \frac{1}{K_n-2} \sum_{m=1}^{K_n-2} V_m^f}_{(T2)} + \underbrace{\frac{1}{K_n} \sum_{j=T_x^{K_n-1}}^n f(X_j)}_{(T3)}. \quad (33)$$

We remark that if $w \in \mathcal{E}_1 \cap \mathcal{E}_2 \cap \mathcal{E}_3 \cap \mathcal{E}_4$, $\lim_{n \rightarrow \infty} K_n(\omega) = +\infty$. Therefore, it's clear that $\lim_{n \rightarrow \infty} (T1)(\omega) = 0$, because $\omega \in \mathcal{E}_4$ implies $T_x(\omega) < +\infty$, and

$$\lim_{n \rightarrow \infty} (T2)(\omega) = \lim_{n \rightarrow \infty} \left(1 - \frac{2}{K_n(\omega)}\right) \frac{1}{K_n(\omega)-2} \sum_{m=1}^{K_n(\omega)-2} V_m^f(\omega) = \mathbb{E}_\pi \left[V_1^f \right], \quad (34)$$

since $\omega \in \mathcal{E}_2$. On the other hand,

$$\begin{aligned} |(T3)(\omega)| &\leq \frac{1}{K_n(\omega)} \sum_{j=T_x^{K_n(\omega)-1}(\omega)}^n |f(X_j(\omega))| \\ &\stackrel{(a')}{\leq} \frac{1}{K_n(\omega)} \sum_{j=T_x^{K_n(\omega)-1}(\omega)}^{T_x^{K_n(\omega)}(\omega)} |f(X_j(\omega))| \\ &= \frac{1}{K_n(\omega)} V_{K_n(\omega)-1}^{|f|}(\omega) \\ &\leq \frac{1}{K_n(\omega)} \max \left\{ V_m^{|f|} : m \in [K_n(\omega)] \right\} \\ &\stackrel{(b')}{\rightarrow} 0, \end{aligned} \quad (35)$$

where the step (a') holds since $n \leq T_x^{K_n(\omega)}(\omega)$, and the step (b') is owing to the fact $\omega \in \mathcal{E}_3$ together with $\lim_{n \rightarrow \infty} K_n(\omega) = +\infty$. Combining the above pieces (34) and (35) together with the fact $\lim_{n \rightarrow \infty} (T1)(\omega) = 0$ deduces

$$\lim_{n \rightarrow \infty} \frac{1}{K_n(\omega)} \sum_{j=1}^n f(X_j(\omega)) = 0 + \mathbb{E}_\pi [V_1^f] + 0 = \mathbb{E}_\pi [V_1^f]$$

for all $w \in \mathcal{E}_1 \cap \mathcal{E}_2 \cap \mathcal{E}_3 \cap \mathcal{E}_4$. Hence, for any $w \in \mathcal{E}_1 \cap \mathcal{E}_2 \cap \mathcal{E}_3 \cap \mathcal{E}_4$,

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{j=1}^n f(X_j(\omega)) &= \lim_{n \rightarrow \infty} \frac{K_n(\omega)}{n} \cdot \frac{1}{K_n(\omega)} \sum_{j=1}^n f(X_j(\omega)) \\ &\stackrel{(c')}{=} \frac{\mathbb{E}_\pi [V_1^f]}{\mathbb{E}_x [T_x]} \\ &= \pi(f), \end{aligned} \tag{36}$$

where the step (c') follows from the fact $\omega \in \mathcal{E}_1$. So, (32) finishes the proof of the statement (iii).

References

- [1] Rick Durrett. *Probability: theory and examples*, volume 49. Cambridge university press, 2019.